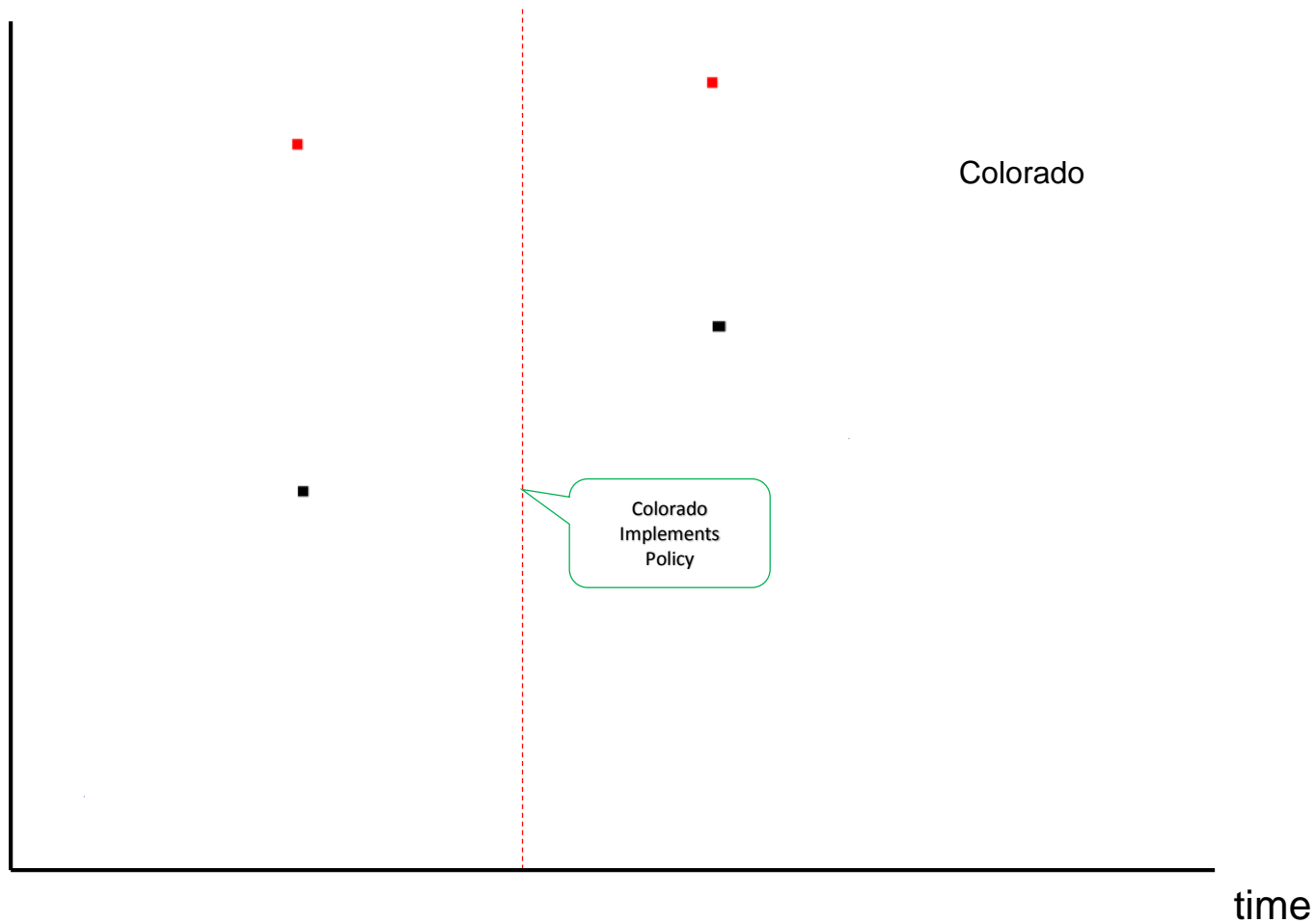
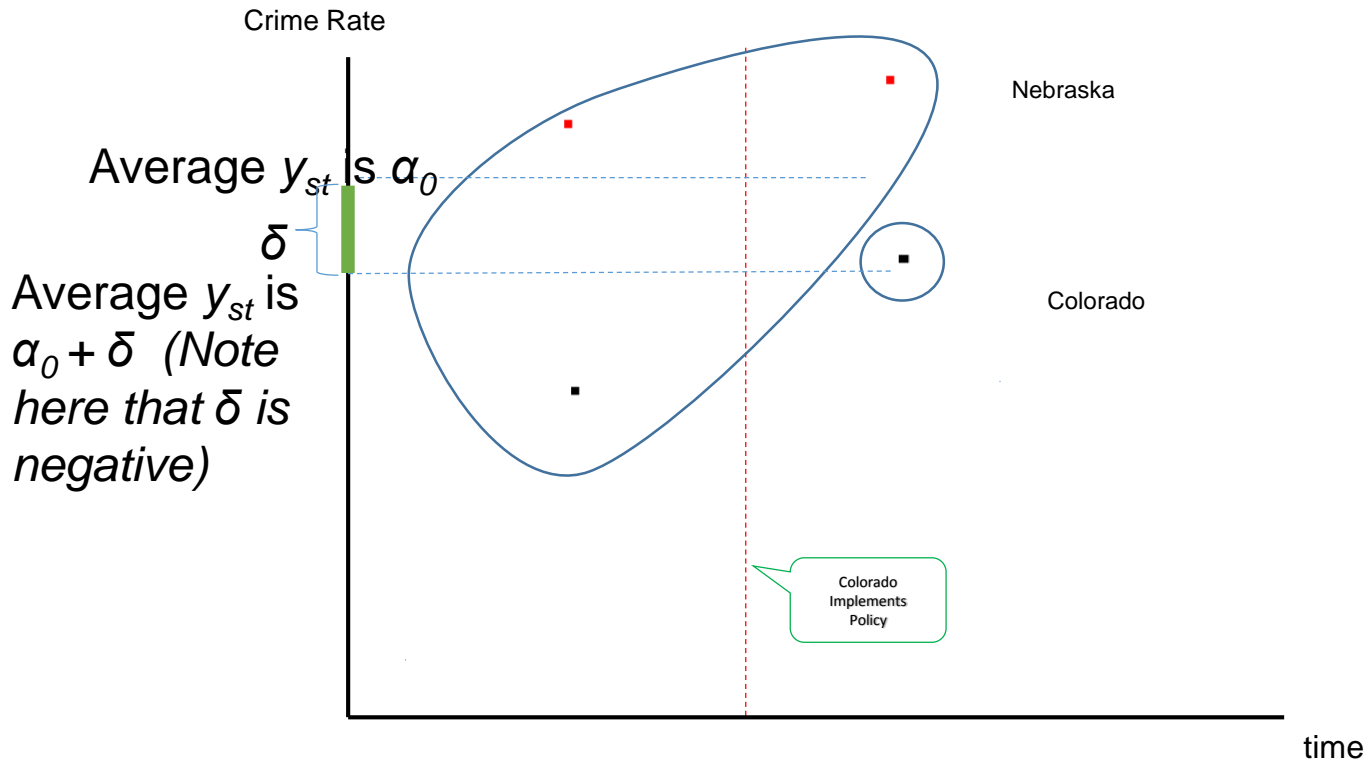


Traffic Fatality Rate



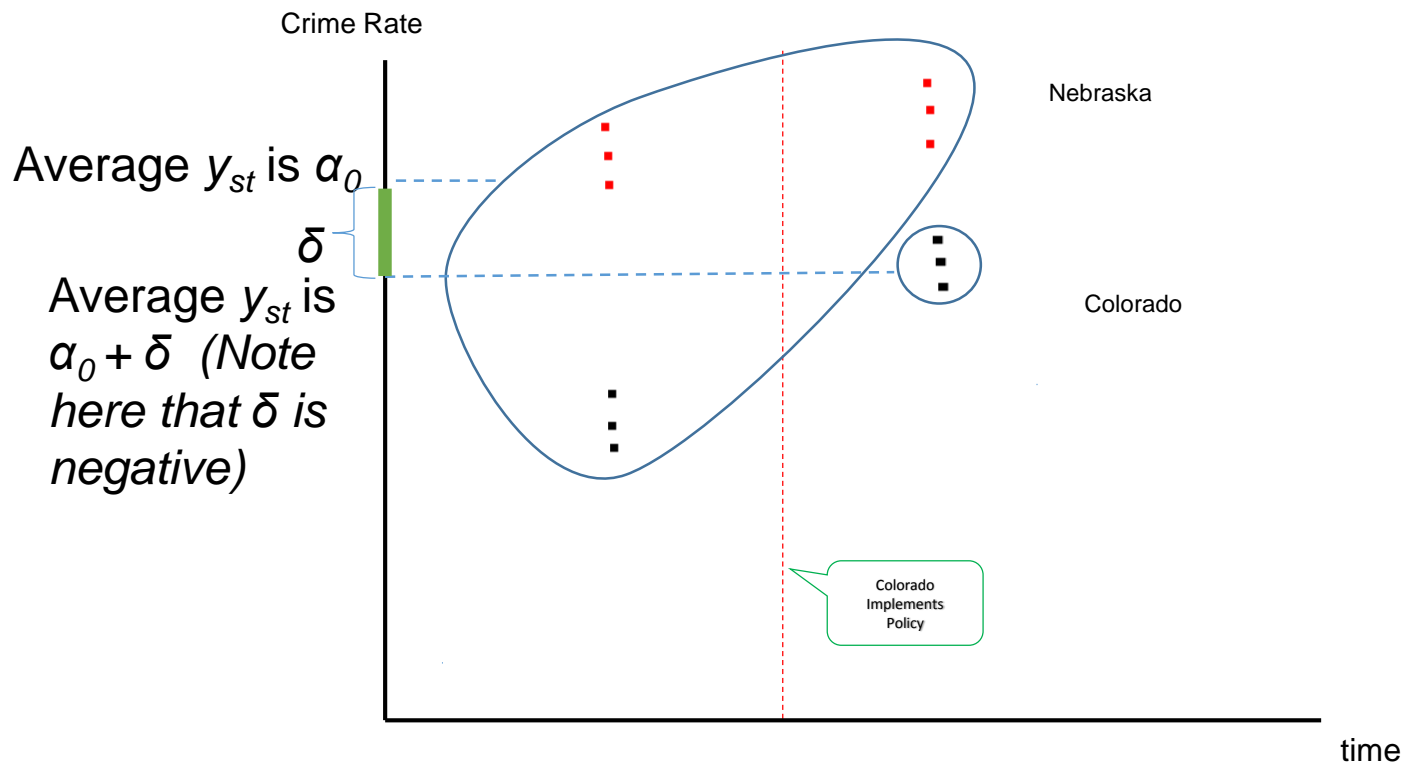
Difference in Difference Model with two states and two time periods. The estimating equation is:

$$y_{st} = \alpha_0 + \delta Treatment_s * Post_t + \epsilon_{st}$$



Suppose the estimating equation is:

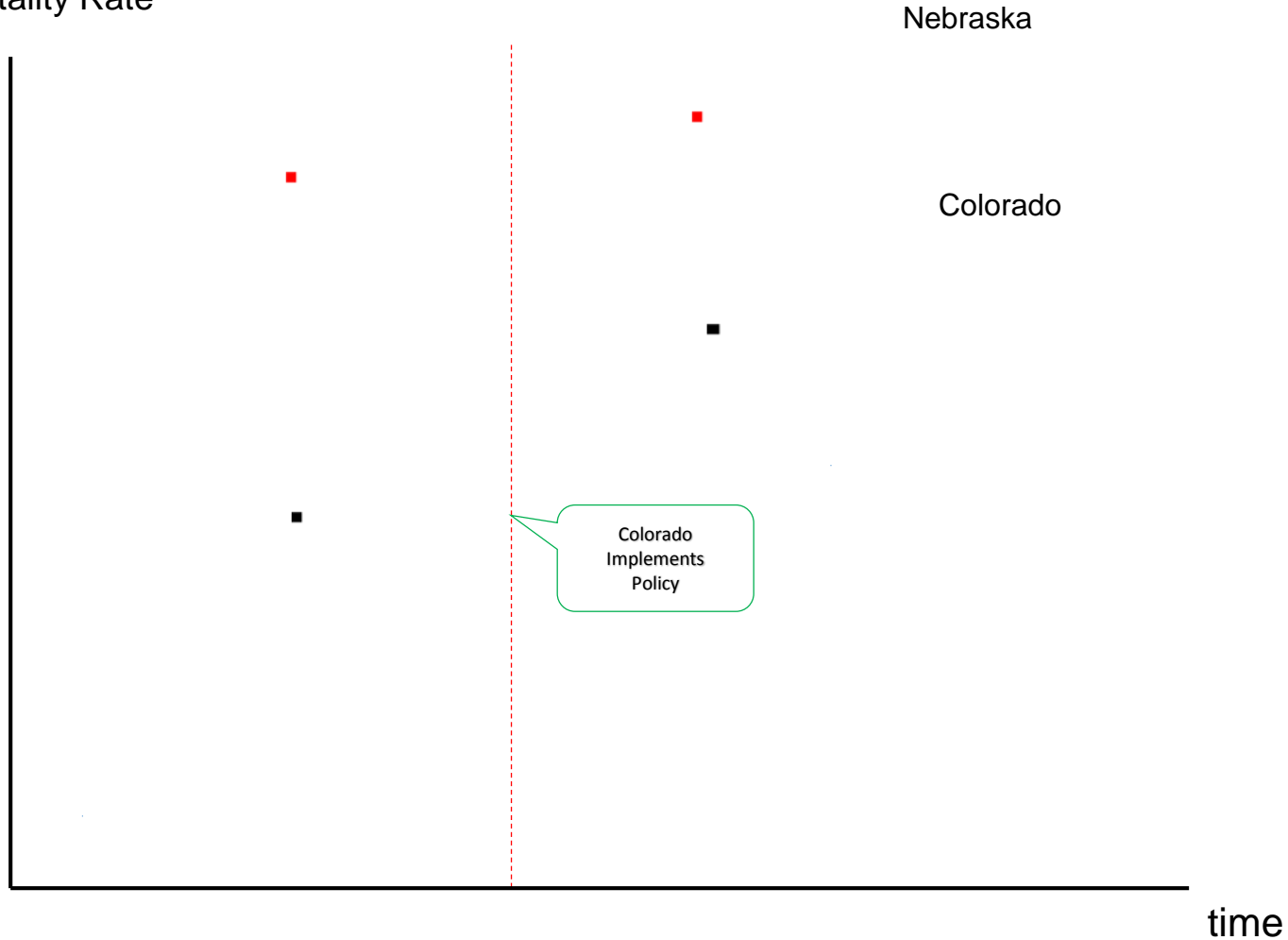
$$y_{st} = \alpha_0 + \delta \text{Treated}_{st} + \varepsilon_{st}$$



Multiple counties in the state. Suppose the estimating equation is:

$$y_{st} = \alpha_0 + \delta Treated_{st} + \varepsilon_{st}$$

Traffic Fatality Rate



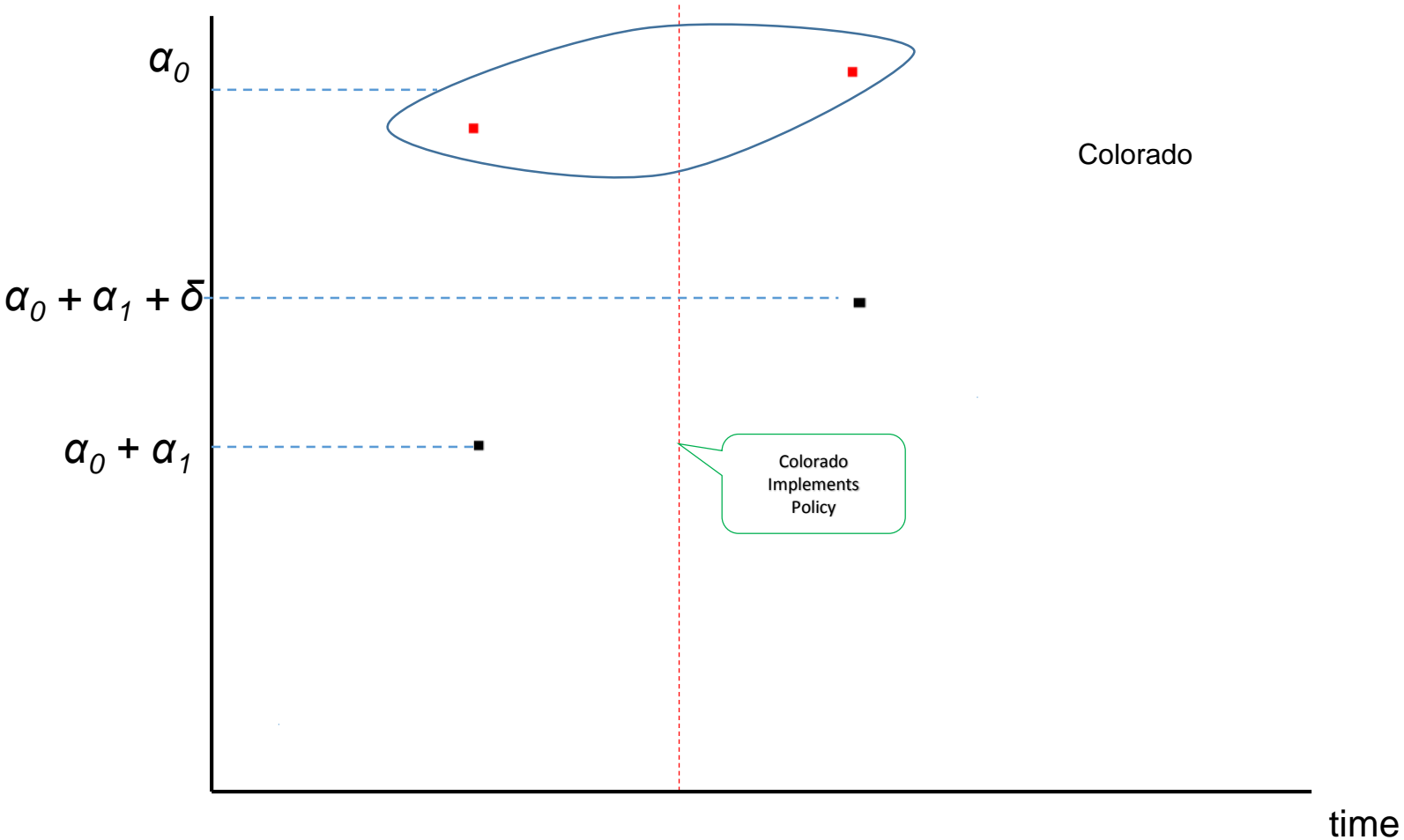
Difference in Difference Model with two states and two time periods. The estimating equation is:

$$y_{st} = \alpha_0 + \alpha_1 Treatment_s + \delta Treatment_s * Post_t + \epsilon_{st}$$

Traffic Fatality Rate

Nebraska

Colorado



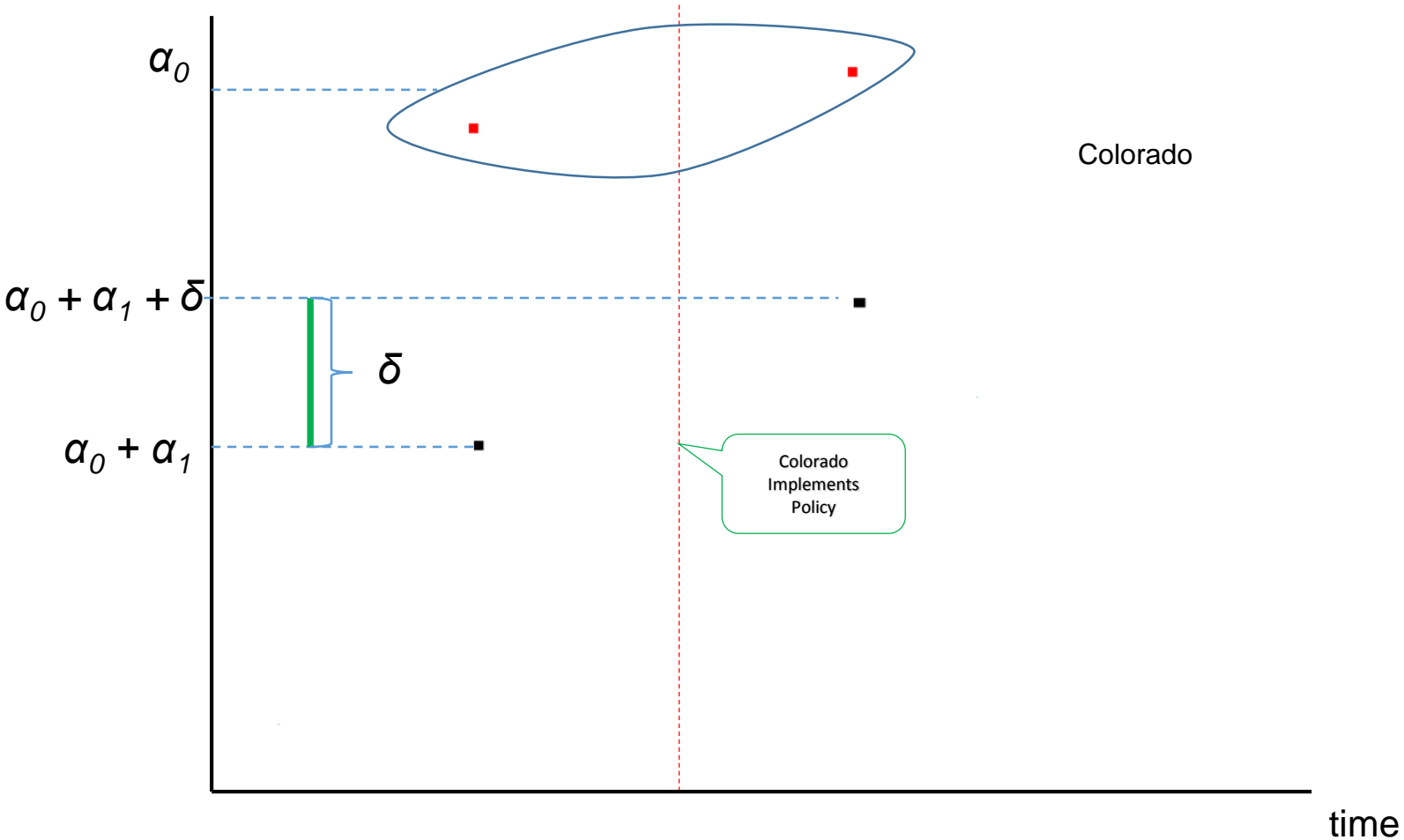
Difference in Difference Model with two states and two time periods. The estimating equation is:

$$y_{st} = \alpha_0 + \alpha_1 Treatment_s + \delta Treatment_s * Post_t + \epsilon_{st}$$

Traffic Fatality Rate

Nebraska

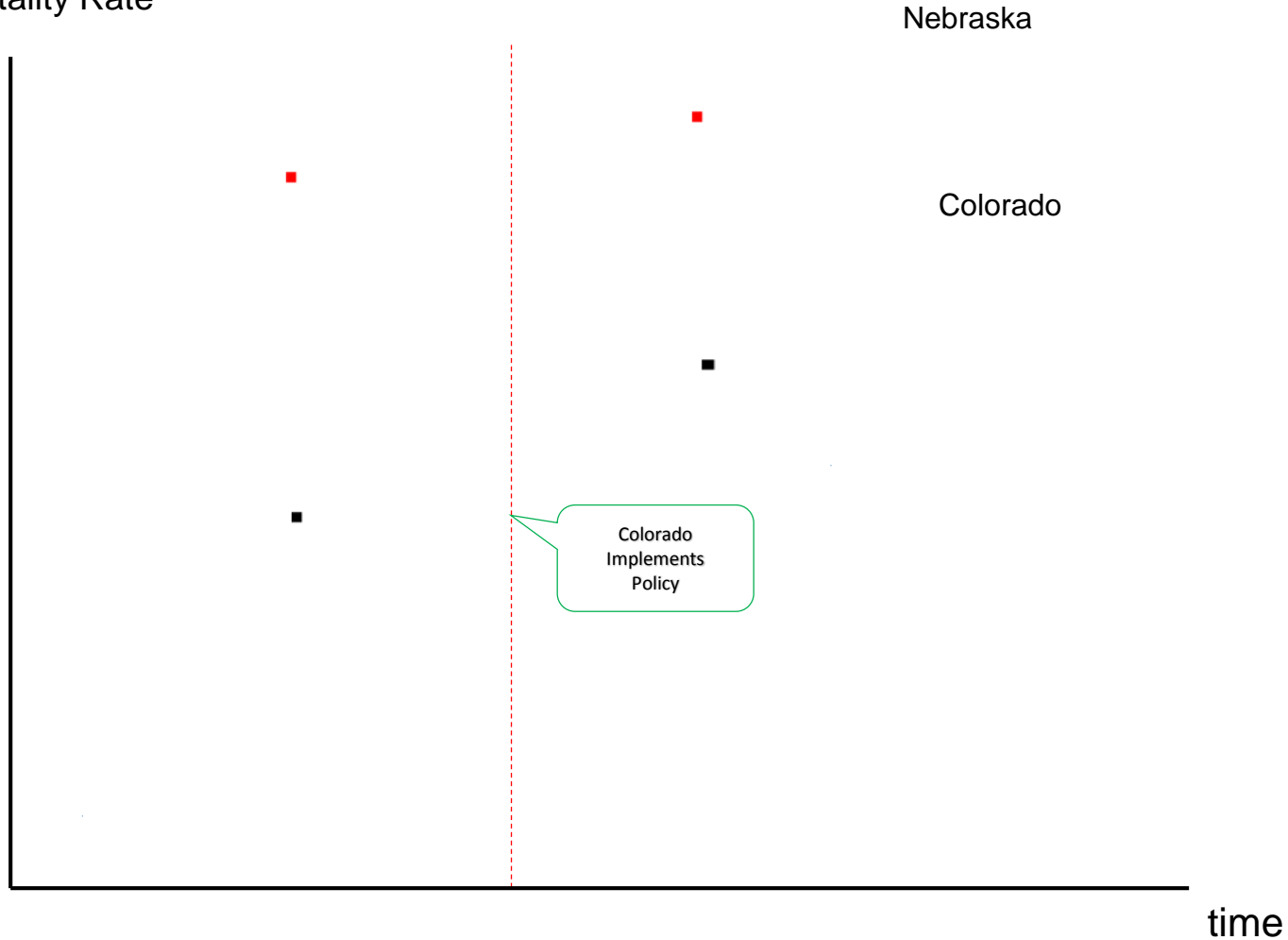
Colorado



Difference in Difference Model with two states and two time periods. The estimating equation is:

$$y_{st} = \alpha_0 + \alpha_1 Treatment_s + \delta Treatment_s * Post_t + \epsilon_{st}$$

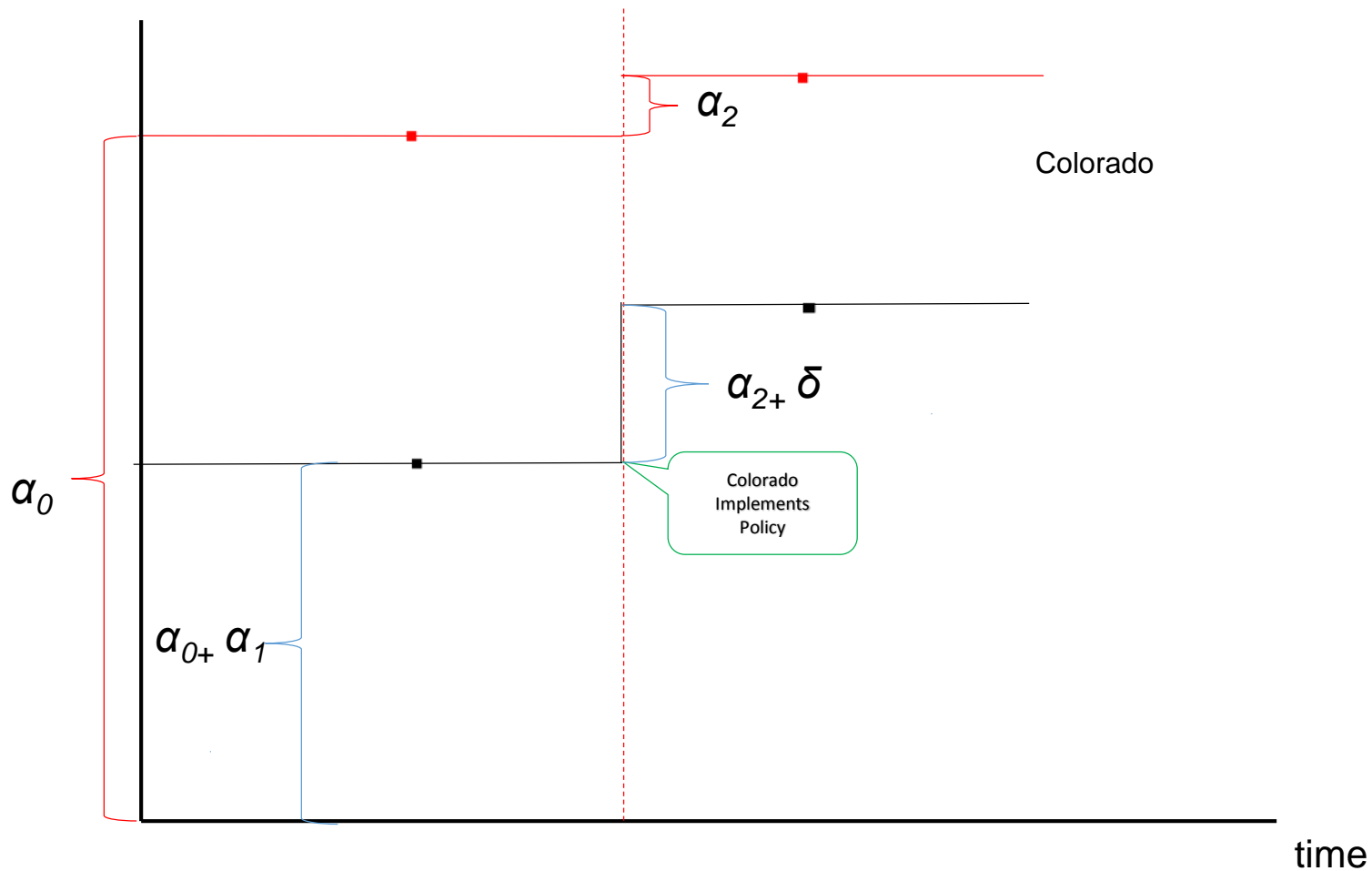
Traffic Fatality Rate



Difference in Difference Model with two states and two time periods. The estimating equation is:

$$y_{st} = \alpha_0 + \alpha_1 Treatment_s + \alpha_2 Post_t + \delta Treatment_s * Post_t + \epsilon_{st}$$

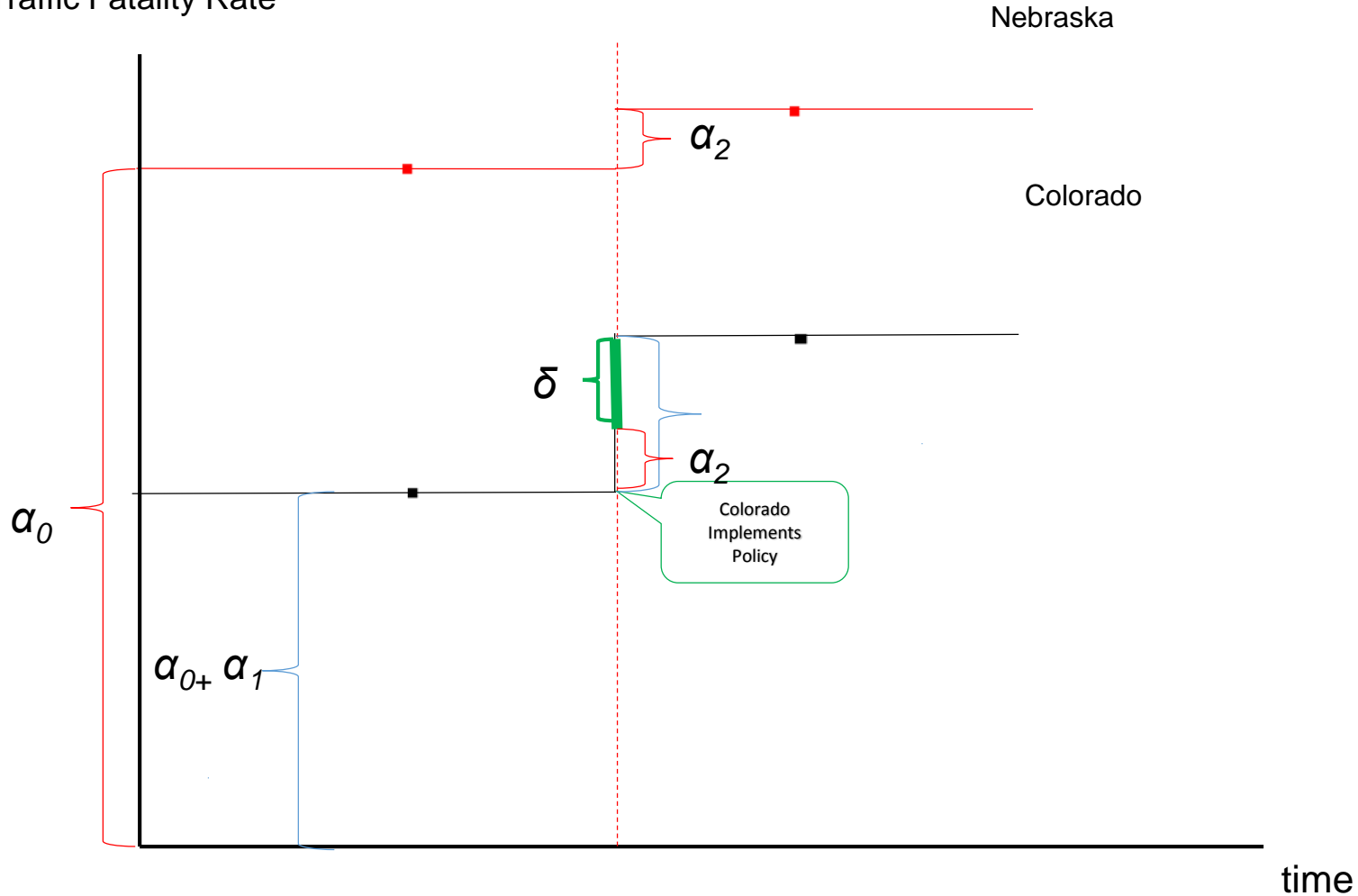
Traffic Fatality Rate



Difference in Difference Model with two states and two time periods. The estimating equation is:

$$y_{st} = \alpha_0 + \alpha_1 Treatment_s + \alpha_2 Post_t + \delta Treatment_s * Post_t + \epsilon_{st}$$

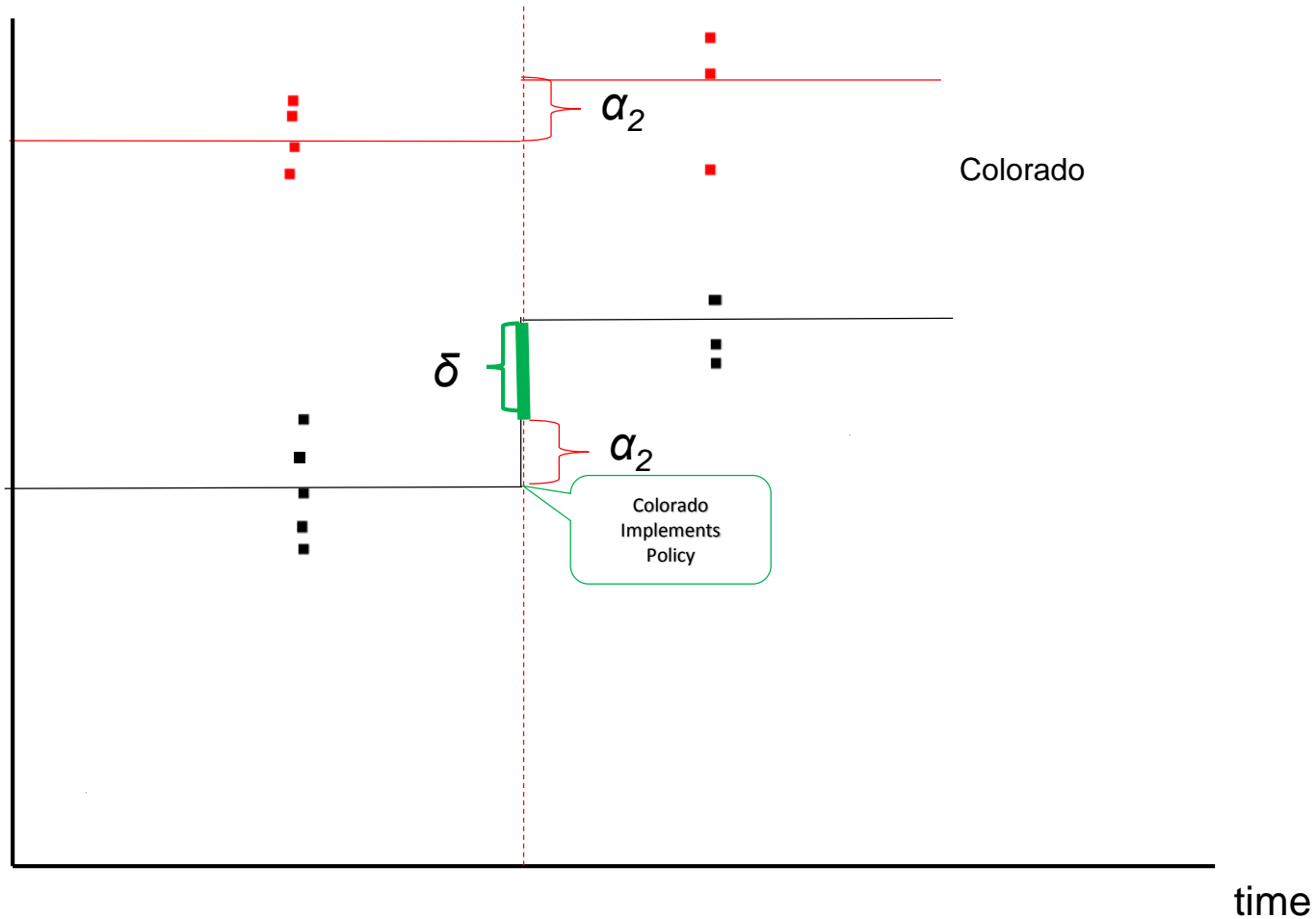
Traffic Fatality Rate



Difference in Difference Model with two states and two time periods. The estimating equation is:

$$y_{st} = \alpha_0 + \alpha_1 Treatment_s + \alpha_2 Post_t + \delta Treatment_s * Post_t + \epsilon_{st}$$

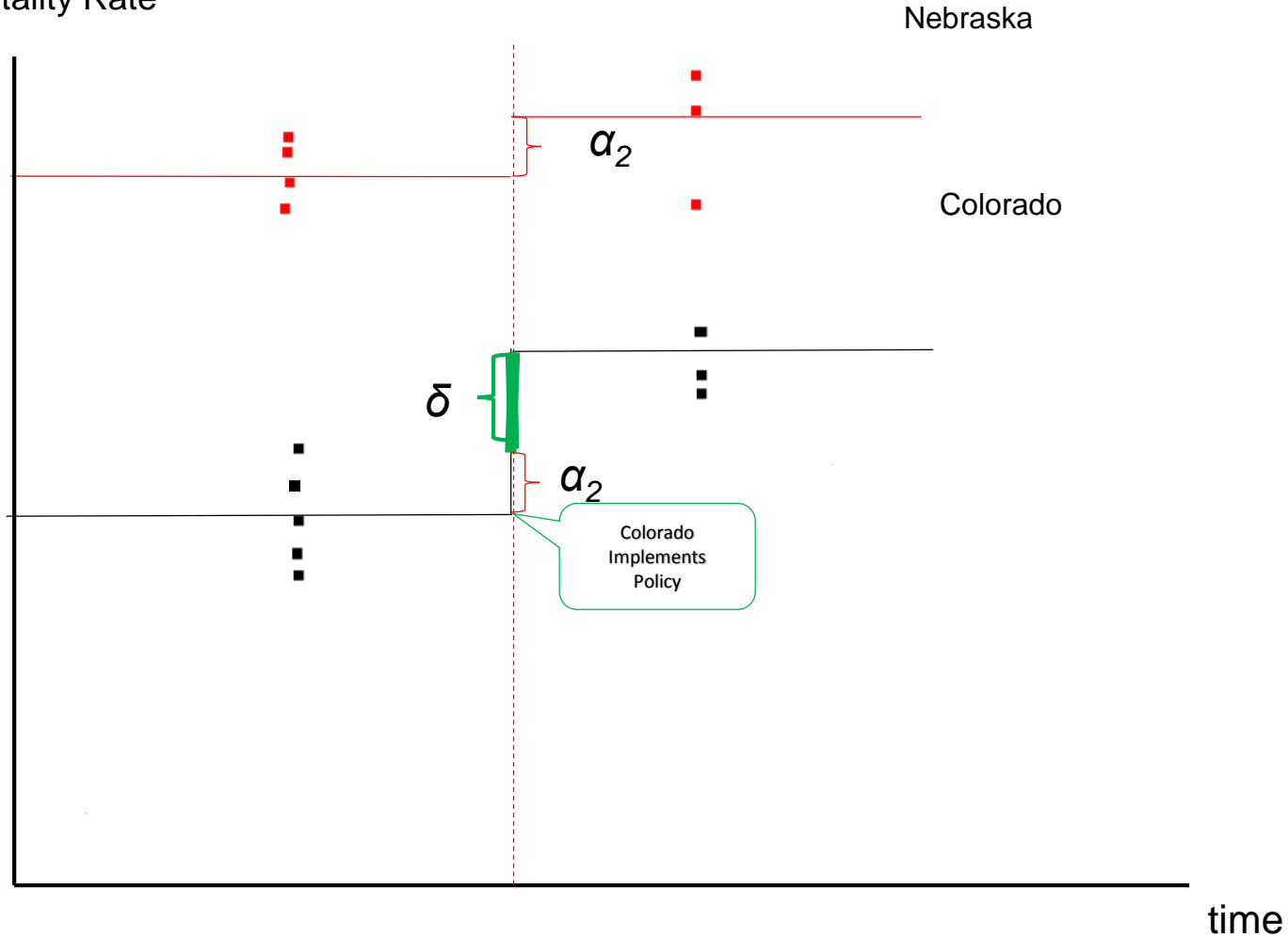
Traffic Fatality Rate



Difference in Difference Model with two states and two time periods. Multiple counties. The estimating equation is:

$$y_{st} = \alpha_0 + \alpha_1 Treatment_s + \alpha_2 Post_t + \delta Treatment_s * Post_t + \epsilon_{st}$$

Traffic Fatality Rate

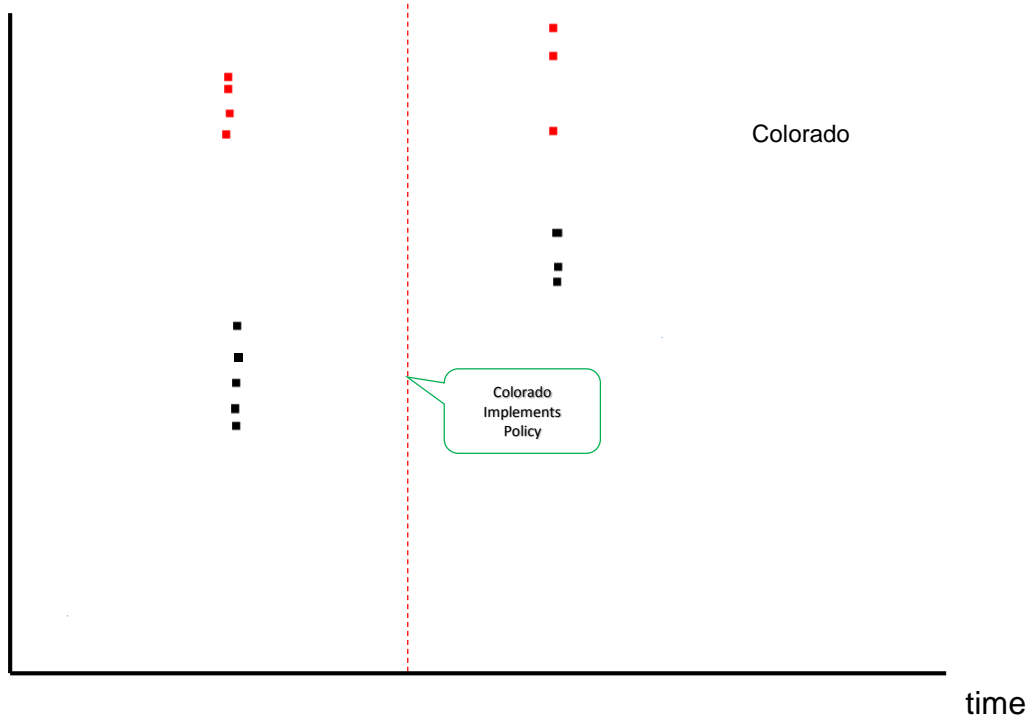


This estimating equation is exactly the same as:

$$y_{ist} = \alpha_0 + \alpha_1 dCO_s + \alpha_2 d2_t + + \delta Policy_{st} + \varepsilon_{ist}$$

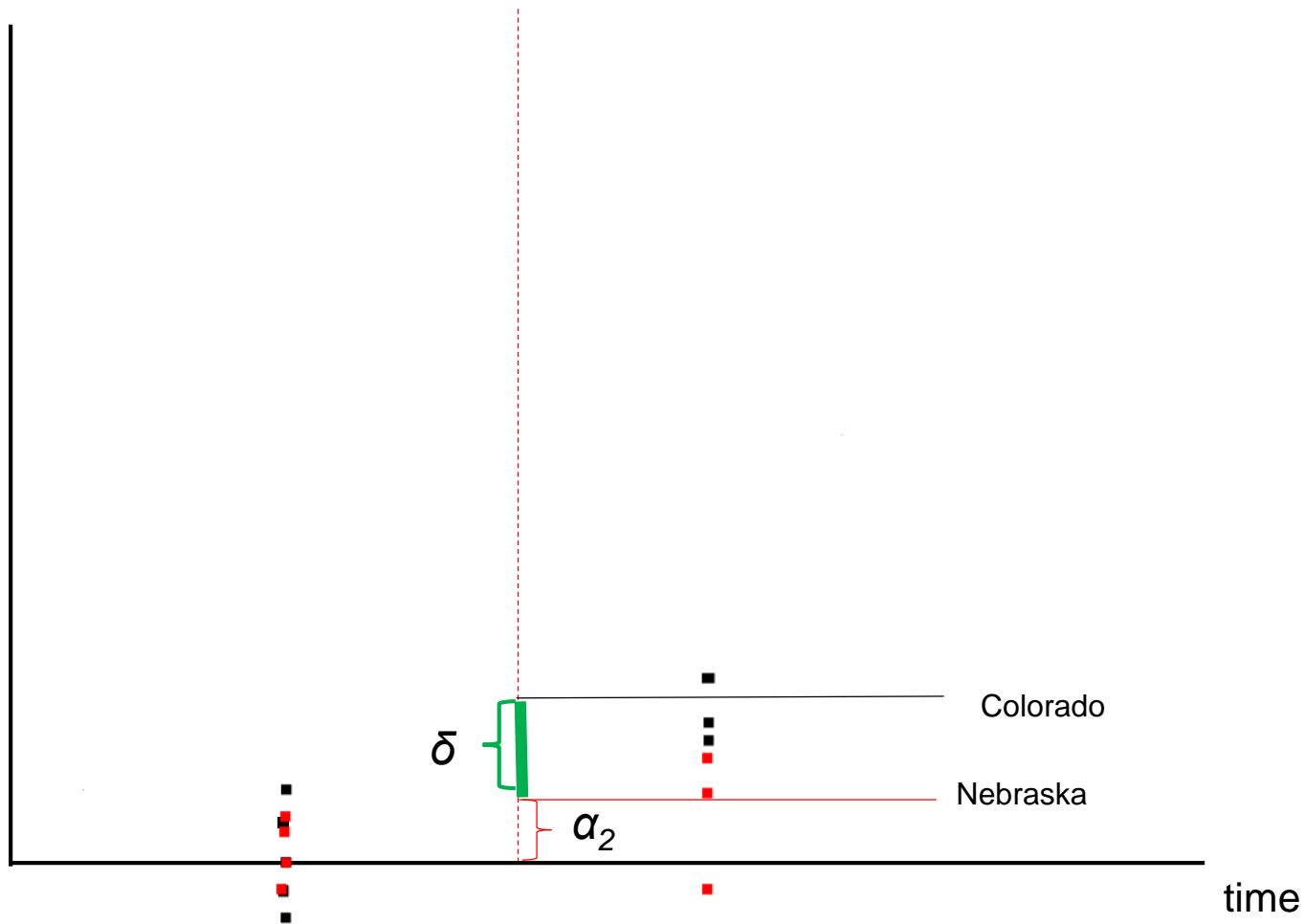
dCO is dummy for CO; $d2$ is dummy for year 2; $Policy$ is dummy =1 in CO after policy is passed

Traffic Fatality Rate



$$y_{ist} - \bar{y}_s = \alpha_2 d2_t + \delta Policy_{st} + \varepsilon_{ist} - \bar{\varepsilon}_s$$

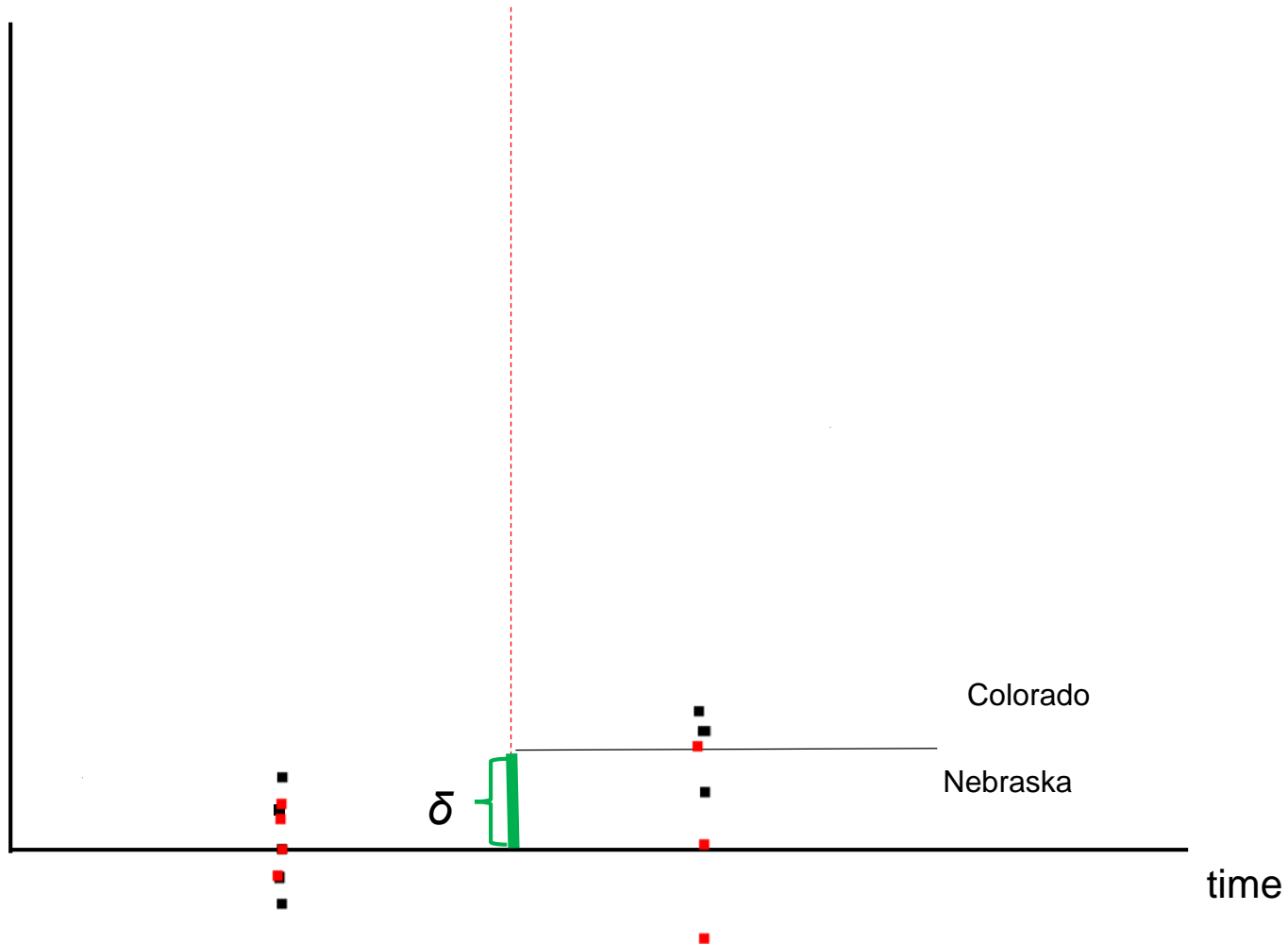
Traffic Fatality Rate



The estimate of δ is exactly the same as obtained by subtracting the mean for each state:

$$y_{ist} - \bar{y}_s = \alpha_2 d_{2t} + \delta Policy_{st} + \varepsilon_{ist} - \bar{\varepsilon}_s$$

Traffic Fatality Rate



And δ is exactly the same after subtracting the mean for each time period:

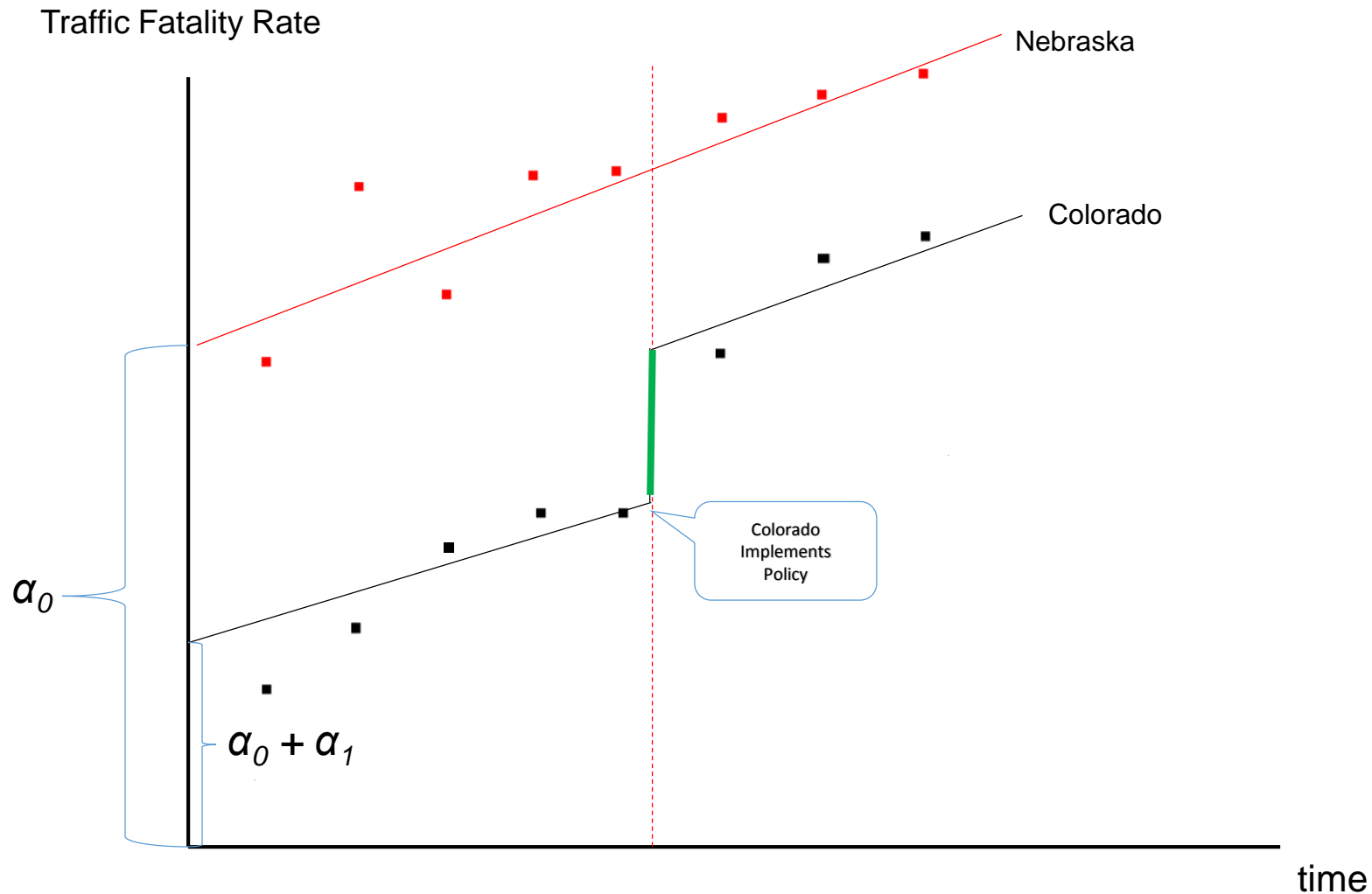
$$y_{ist} - \bar{y}_s - \bar{y}_t = \delta Policy_{st} + \varepsilon_{ist} - \bar{\varepsilon}_s - \bar{\varepsilon}_t$$

When we have lots of states and years, an author typically writes

$$y_{st} = \beta_0 + \delta Policy_{st} + v_s + z_t + \varepsilon_{st}$$

And then the author might say that the equation is estimated including state and year fixed effects

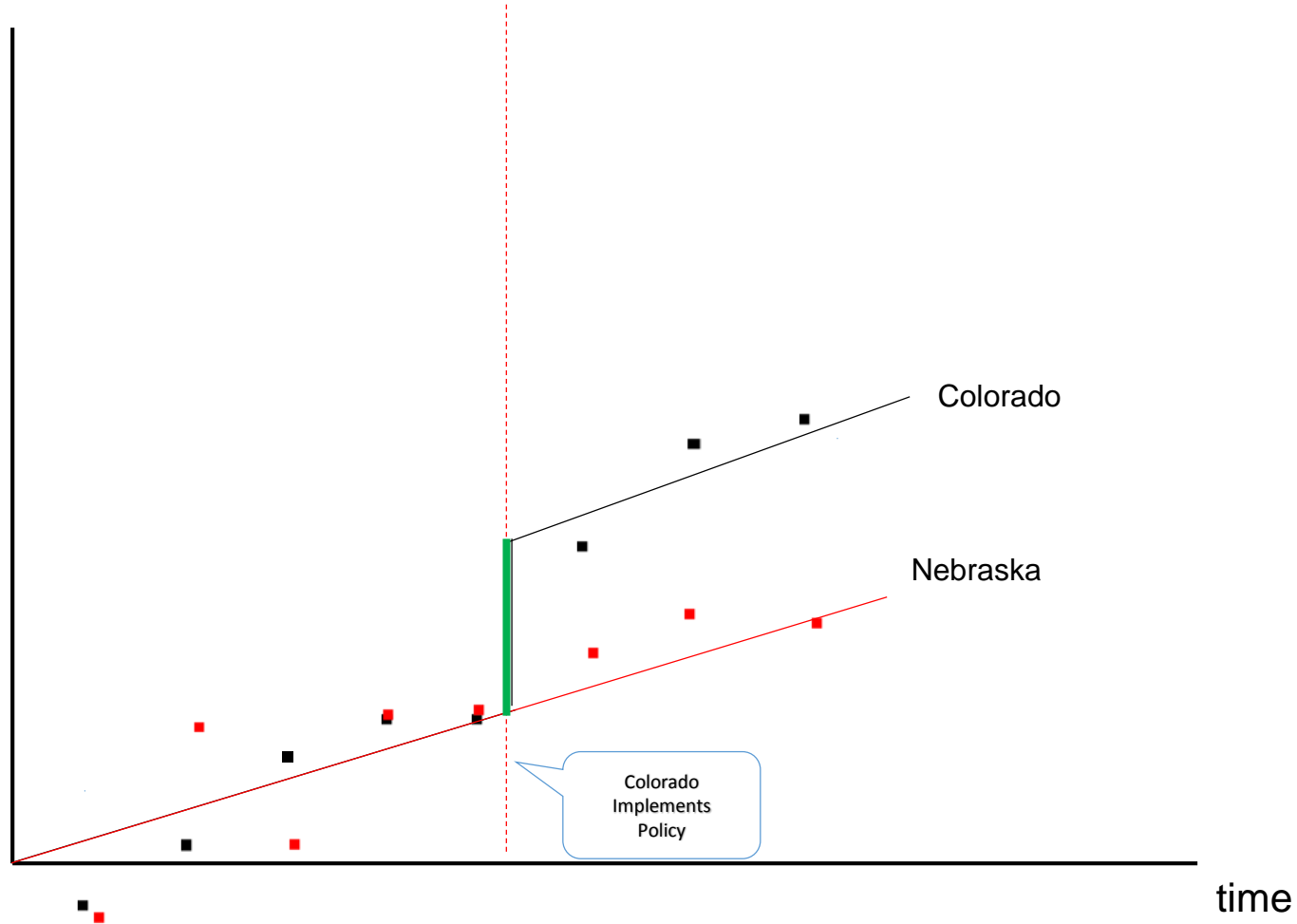
Start with state fixed effects, common time trend



If we're using state-level data, then each state contributes one observation per year. The estimating equation with a common time trend is

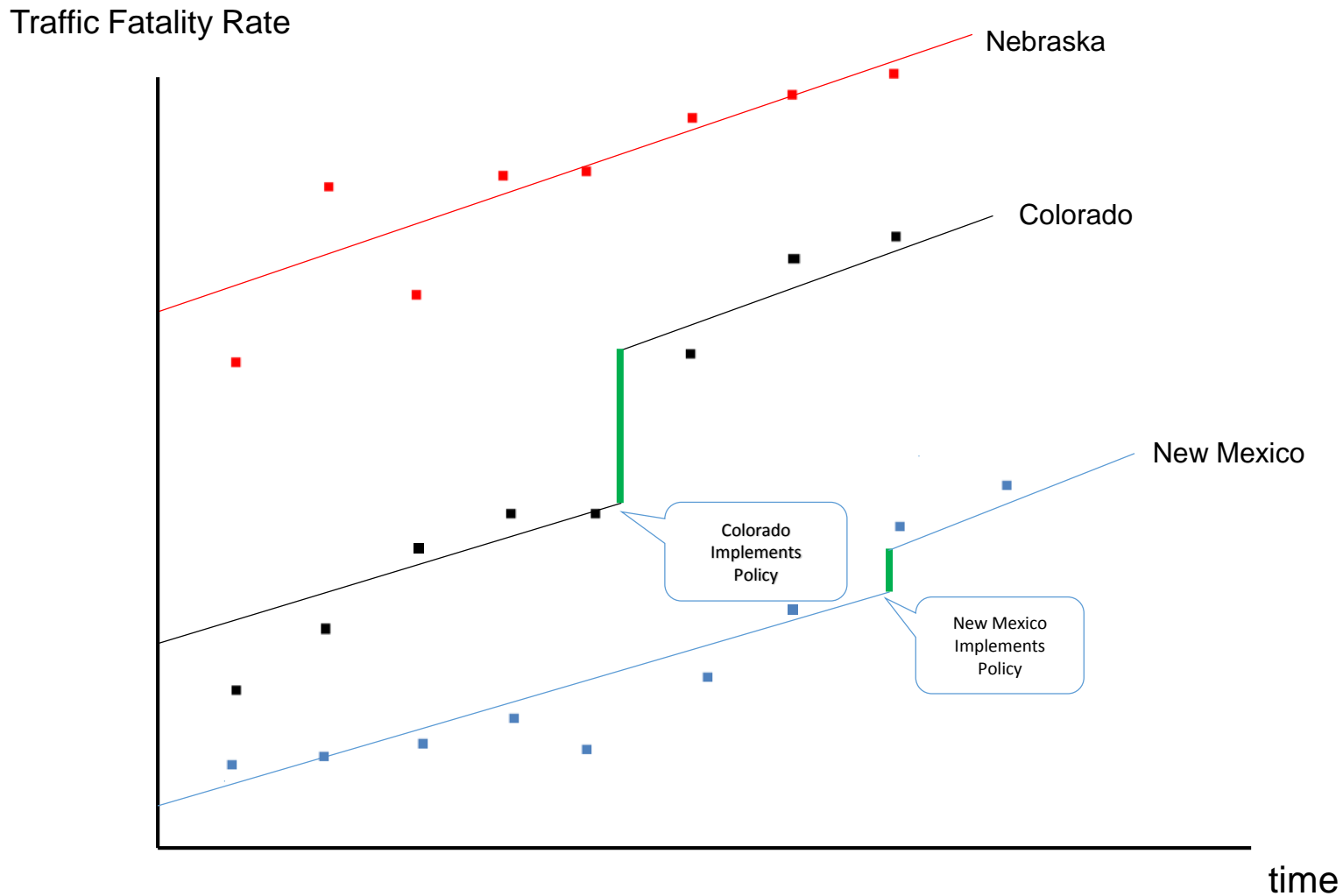
$$y_{ist} = \alpha_0 + \alpha_1 dCO_s + \alpha_2 t + \delta Policy_{st} + \varepsilon_{ist}$$

Traffic Fatality Rate



Remove state specific means $\rightarrow \delta$ will still be the vertical change

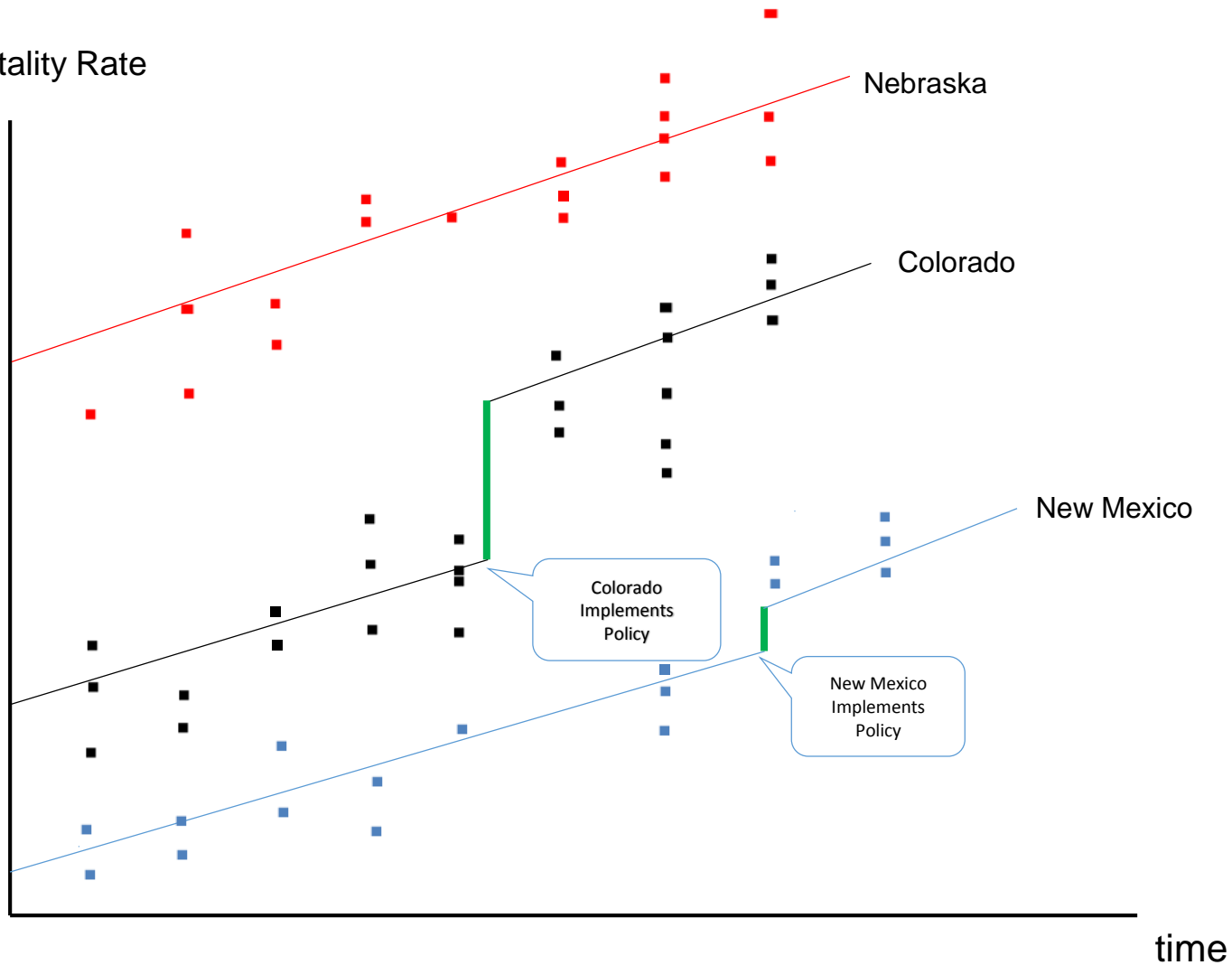
$$y_{ist} - \bar{y}_s = \alpha_2 t + \delta Policy_{st} + \varepsilon_{ist} - \bar{\varepsilon}_s$$



With multiple states δ will be the average vertical change across the states

$$y_{st} = \alpha_0 + \alpha_2 t + \delta Policy_{st} + v_s + \varepsilon_{st}$$

Traffic Fatality Rate



If we're using repeated cross-sectional data at the individual level, then each state contributes multiple observations per year. The estimating equation is:

$$y_{ist} = \alpha_0 + \delta Policy_{st} + \alpha_2 t + v_s + \varepsilon_{ist}$$

Pooled model: $MR_{it} = \beta_0 + \beta_1 unem_{it} + e_{it}$

id	state	year	murder rate			unem			
19	LA	87	11.1			12			
19	LA	90	17.2			6.2			
19	LA	93	20.3			7.4			
5	CA	87	10.6			5.8			
5	CA	90	11.9			5.6			
5	CA	93	13.1			9.2			

$$\text{FD model: } \Delta \text{MR}_{it} = \beta_1 \Delta \text{unem}_{it} + \Delta e_{it}$$

id	state	year	murder rate	ΔMR	unem	ΔUN
19	LA	87	11.1	--	12	--
19	LA	90	17.2	6.1	6.2	-5.8
19	LA	93	20.3	3.1	7.4	1.2
5	CA	87	10.6	--	5.8	--
5	CA	90	11.9	1.3	5.6	-.02
5	CA	93	13.1	1.2	9.2	3.6

$$\text{FE model: } MR_{it} - \overline{MR}_i = \beta_1 \text{unem}_{it} - \overline{\text{unem}}_i + e_{it}$$

			murder		MR-			UN-
id	state	year	rate	Mean	mean	unem	mean	mean
19	LA	87	11.1	16.2	-5.1	12	8.5	3.5
19	LA	90	17.2	16.2	1	6.2	8.5	-2.3
19	LA	93	20.3	16.2	4.1	7.4	8.5	-1.1
5	CA	87	10.6	11.9	-1.3	5.8	6.9	-1.1
5	CA	90	11.9	11.9	0	5.6	6.9	-1.3
5	CA	93	13.1	11.9	1.2	9.2	6.9	2.3

The estimated coefficient on unemployment from these two models will be the same if there are only 2 years.

Otherwise, they are both consistent estimators of β_1 but their exact values will differ

$$\text{FE model: } MR_{it} = \alpha_0 + \beta_1 \text{unem}_{it} + \alpha_1 \text{State}_i + e_{it}$$

id	state	year	murder		unem	State
			rate			
19	LA	87	11.1		12	1
19	LA	90	17.2		6.2	1
19	LA	93	20.3		7.4	1
5	CA	87	10.6		5.8	0
5	CA	90	11.9		5.6	0
5	CA	93	13.1		9.2	0

$$\text{FE model: } MR_{it} = \alpha_0 + \beta_1 \text{unem}_{it} + \alpha_1 \text{State}_i + e_{it}$$

id	state	year	murder		unem	State
			rate			
19	LA	87	11.1		12	1
19	LA	90	17.2		6.2	1
19	LA	93	20.3		7.4	1
5	CA	87	10.6		5.8	0
5	CA	90	11.9		5.6	0
5	CA	93	13.1		9.2	0

The constant term in this model and the previous version of the FE model will be different (how?)

The estimated coefficient on unemployment WILL BE EXACTLY THE SAME

CAN LABOR REGULATION HINDER ECONOMIC PERFORMANCE?
EVIDENCE FROM INDIA¹
TIMOTHY BESLEY AND ROBIN BURGESS

This paper investigates whether the industrial relations climate in Indian states has affected the pattern of manufacturing growth in the period 1958-92. We show that states which ammended the Industrial Disputes Act in a pro-worker direction experienced lowered output, employment, investment and productivity in registered or formal manufacturing. In contrast, output in unregistered or informal manufacturing increased. Regulating in a pro-worker direction was also associated with increases in urban poverty. This suggests that attempts to redress the balance of power between capital and labor can end up hurting the poor.

Our econometric analysis is based on panel data regressions of the form:

$$y_{st} = \alpha_s + \beta_t + \mu r_{st-1} + \xi x_{st} + \varepsilon_{st}$$

where y_{st} is a (logged) outcome variable in state s at time t , r_{st} is the regulatory measure (which we lag one period to capture the gap between enactment

and implementation)¹³, x_{st} are other exogenous variables, α_s is a state fixed effect and β_t is a year fixed effect. We cluster our standard errors by state to deal with concerns with serial correlation (Bertrand, Duflo and Mullainathan, [2002]).¹⁴

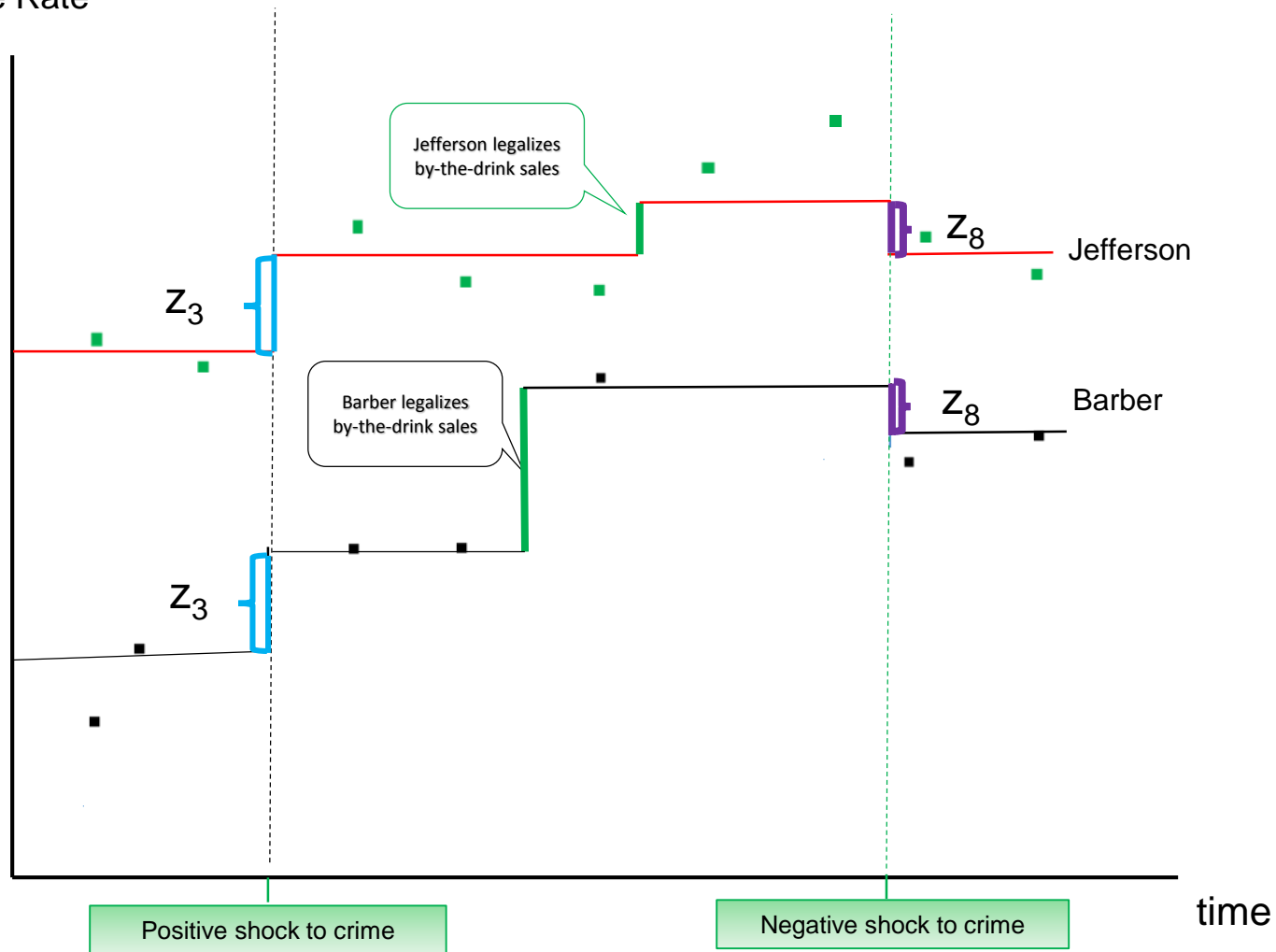
The state fixed effect captures state-specific factors such as culture and geography. The year effects capture common shocks such as central government amendments to the Industrial Disputes Act which took place in 1976 and 1982 (see Fallon [1987] and Fallon and Lucas [1993]) as well as other centrally implemented policies.

Now look at time shocks.

What if all states experience a common shock in a given year?

What if the means varies by time as well as by state?

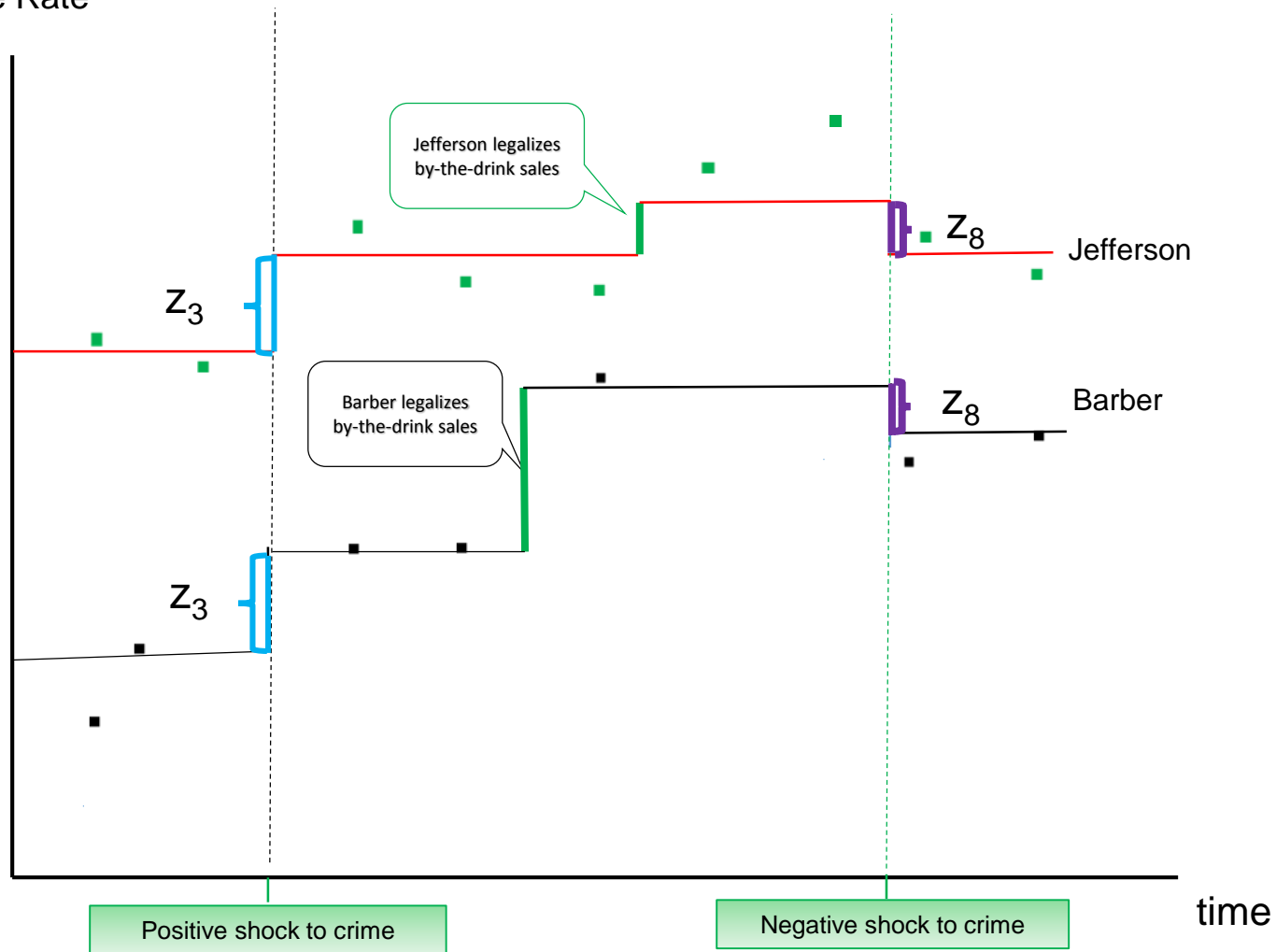
Crime Rate



Common time fixed effects-- δ will still be the average vertical change from before and after the policy

$$Violent\ Crime_{ct} = \pi_0 + \delta Wet\ Law_{ct} + X'_{ct}\pi_2 + v_c + Z_t + \epsilon_{ct}$$

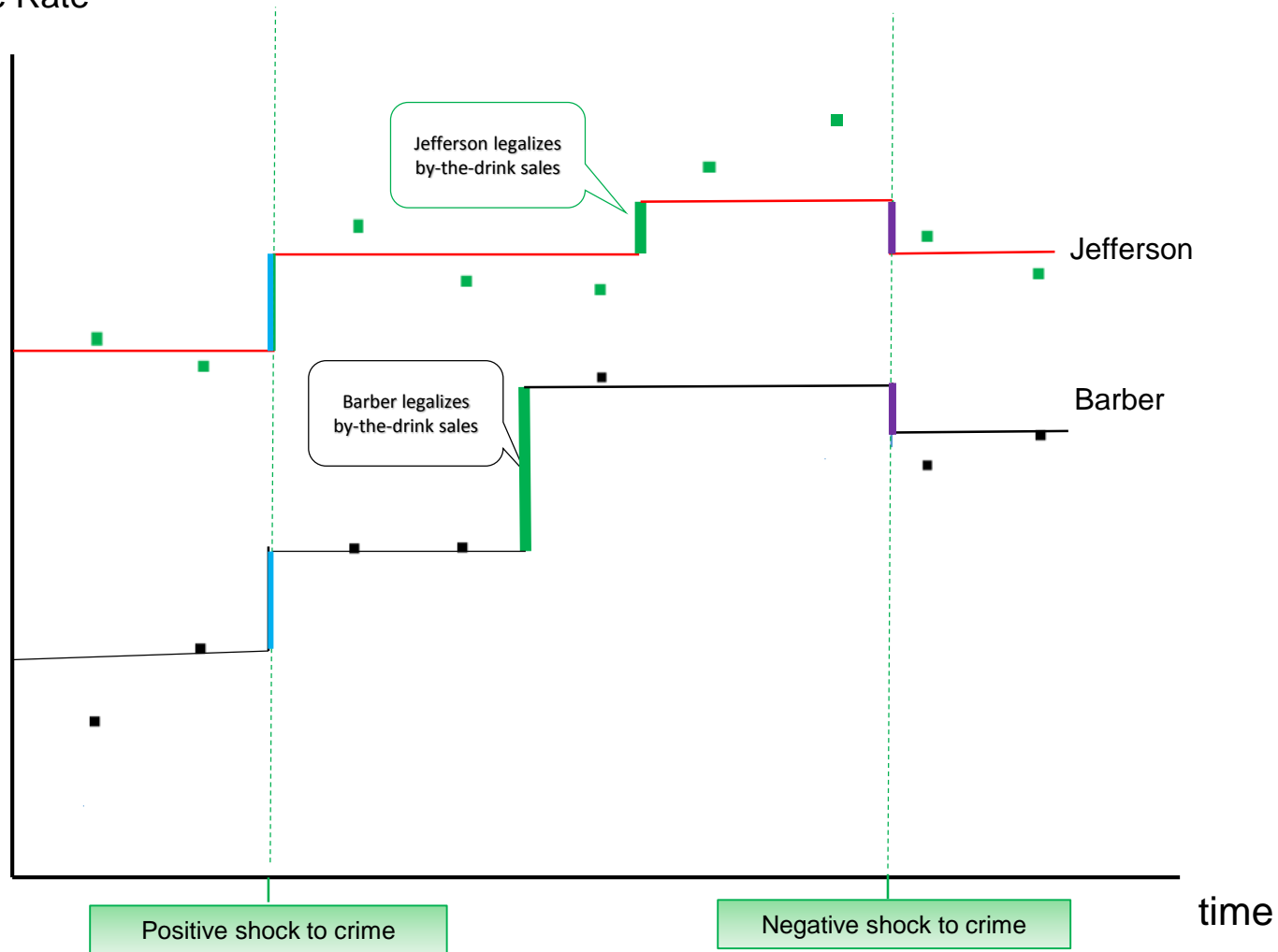
Crime Rate



Common time fixed effects-- δ will still be the average vertical change from before and after the policy

$$Violent\ Crime_{ct} = \pi_0 + \delta Wet\ Law_{ct} + X'_{ct}\pi_2 + v_c + Z_t + \epsilon_{ct}$$

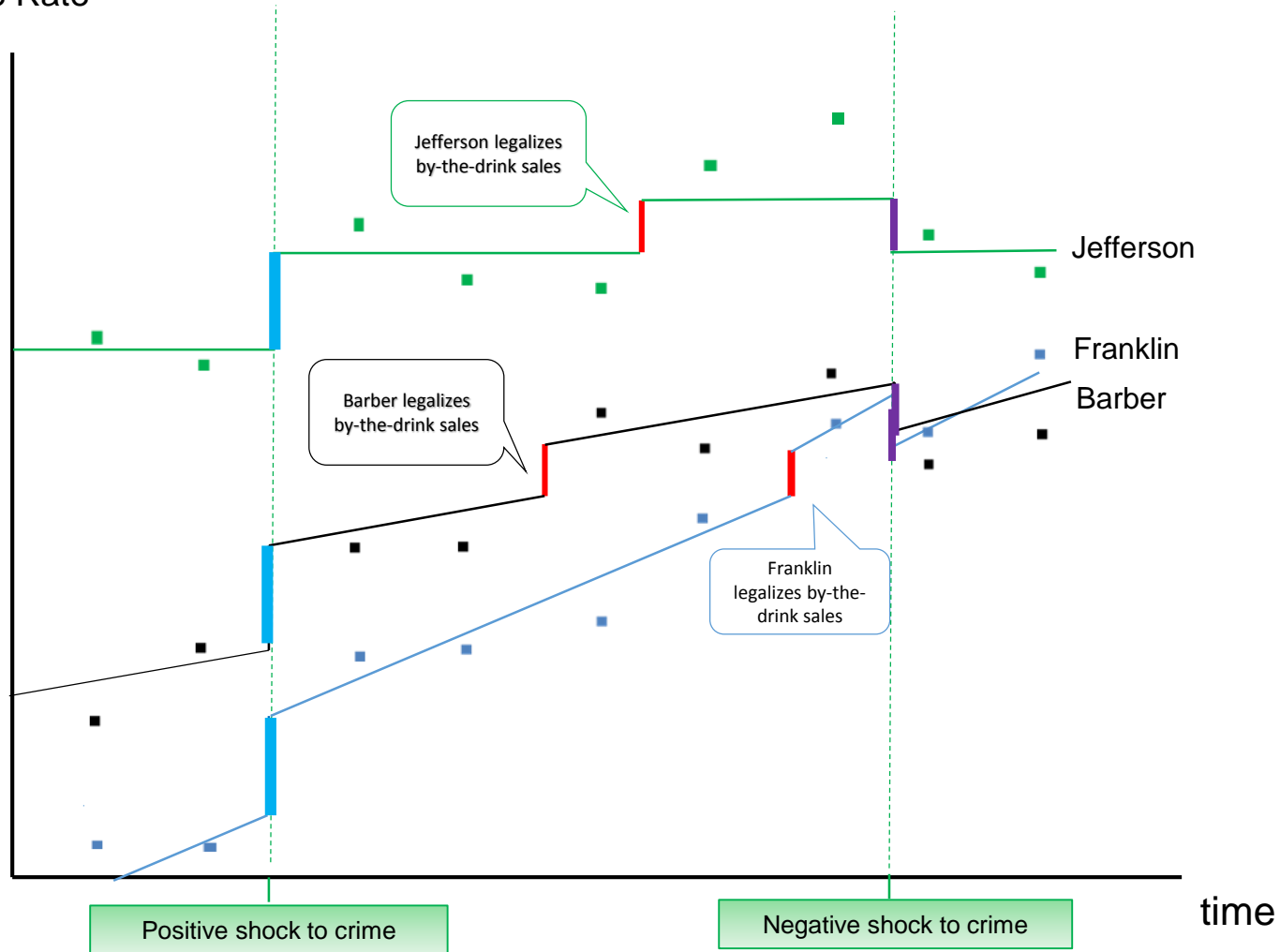
Crime Rate



But note if you look at it, Barber is increasing faster than Jefferson:

$$Violent\ Crime_{ct} = \pi_0 + \pi_1 Wet\ Law_{ct} + X'_{ct}\pi_2 + v_c + Z_t + \epsilon_{ct}$$

Crime Rate



Adding state specific time trends

$$Violent\ Crime_{ct} = \pi_0 + \delta Wet\ Law_{ct} + X_{ct}\pi_2 + v_c + Z_t + \Theta_c \cdot t + \varepsilon_{ct}$$

The fact that our results are not robust to state-specific time trends does raise the question of whether the effects that we are picking up are those due to labor regulations per se or the consequences of a poor climate of labor relations—union power and labor/management hostility—which affect the trend rate of growth within a state. This goes to interpretation of the finding. But either way, the analysis suggests that labor market institutions in India have had an important impact on manufacturing development.

The analysis reinforces the growing sentiment that government regulations in developing countries have not always promoted social welfare. The example that we have studied here is highly specific and it is clear that it cannot be used to promote a generalized pro- or antiregulation stance. Future progress will likely rest on improving our knowledge of specific regulatory policies. Research involving particular country experiences will be an important component of this. Only then can the right balance between the helping and hindering hands of government be found.